

The Joint Meeting of the 2011 Taipei International Statistical Symposium and IASC-CAR (Joint2011)

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A Discussion on Environmental Risks for Breast and Liver Cancer by Analyzing (Age, period)-tabulated Data

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Outline

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Background

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Dietary Soy Isoflavones Inhibit Estrogen Effects in the Postmenopausal Breast

Cancer Res 2006;66:1241-1249.

Charles E. Wood¹, Thomas C. Register¹, Adrian A. Franke², Mary S. Anthony¹, and J. Mark Cline¹

In the presence of estrogen higher doses of dietary soy isoflavones may alter estrogen receptors signaling and induce selective antagonistic effects in the breast.

Background

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High soybean consumption increases risk of liver cancer in women: survey

<http://www.hepcaustralia.com.au/symptoms-news/high-soybean-consumption-increases-risk-of-liver-cancer-in-women-survey>

Women who eat a lot of soybean products are three to four times more likely to develop liver cancer than those who eat only a small amount.

Estrogen heightens the risk of breast cancer, but protects against liver cancer. It is believed that excessive consumption of isoflavones may block this property.

Aim

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Our aim in this study is

to examine whether there exist environmental factors which decrease the breast cancer risk and increase the liver cancer risk for postmenopausal females aged 55 and over based on the analysis of (age, period)-tabulated data.

Procedure

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Our procedure is

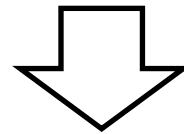
- (1) estimating the environmental risks of breast cancer and liver cancer for female based on (age, period)-tabulated data,
- (2) testing for the significance of the correlation coefficient between environmental risks for breast cancer and liver cancer.

Procedure

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However

Hepatitis B virus infection is the main causes of liver cancer. Liver cancer correlates with the frequency of chronic hepatitis B virus infection.



It is needed to eliminate the effect of infection to hepatitis B virus from environment effects.

Procedure

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Difference between sexes

Isoflavone consumption and subsequent risk of hepatocellular carcinoma in a population-based prospective cohort of Japanese men and women

Int J Cancer. 2009 Apr 1;124(7):1644-9

Kurahashi N, Inoue M, Iwasaki M, Tanaka Y, Mizokami M, Tsugane S; JPHC Study Group

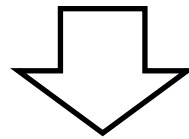
In women, genistein and daidzein were dose-dependently associated with an increased risk of hepatocellular carcinoma (HCC, liver cancer). No association between isoflavones and HCC was observed in men.

Procedure

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Difference between sexes

In women, genistein and daidzein were dose-dependently associated with an increased risk of hepatocellular carcinoma (HCC, liver cancer). No association between isoflavones and HCC was observed in men.



The difference of environmental effect on liver cancer between men and women is effect of isoflavone on breast cancer for women.

Procedure

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Our modified procedure is

- (1) estimating the environmental risks of breast cancer for female and liver cancer for both sexes,
- (2) calculating difference of environmental risks between men and women,
- (3) testing for the significance of the correlation coefficient between environmental effects on breast cancer and **the difference** of environmental effect on liver cancer **between men and women**.

(Age, period)-tabulated data

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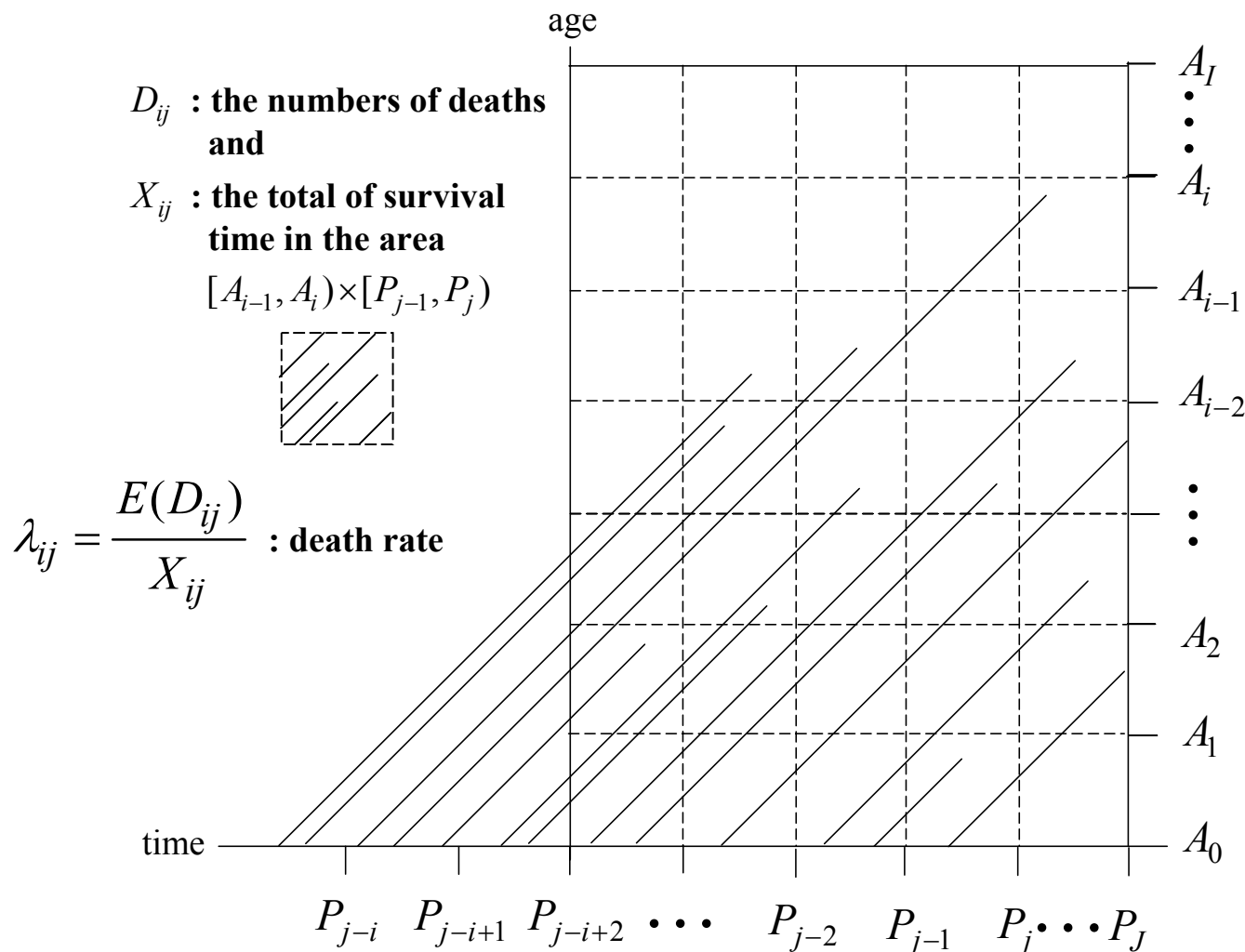
Table 1. Deaths of liver cancer in 100 thousand Japanese males aged 55 and over

age	Period							
	1970-74	1975-79	1980-84	1985-89	1990-94	1995-99	2000-04	2005-09
55~59	14.5	13.7	12.7	13.7	12.5	10.5	8.2	6.3
60~64	22.8	21.1	21.8	23.0	25.0	26.4	18.4	12.7
65~69	35.8	32.9	32.3	35.4	38.4	47.2	40.4	27.0
70~74	50.5	46.9	46.6	48.6	51.9	62.0	64.4	54.5
75~79	66.4	62.1	63.9	62.0	65.6	75.6	78.9	78.1

Deaths of liver and breast cancer in 100 thousand Japanese females aged 55 and over are given in the same formula.

Lexis diagram

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Age period cohort model

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$$\log \lambda_{ij} = \mu + \alpha_i + \beta_j + \gamma_{j-i}$$

α_i : age effect

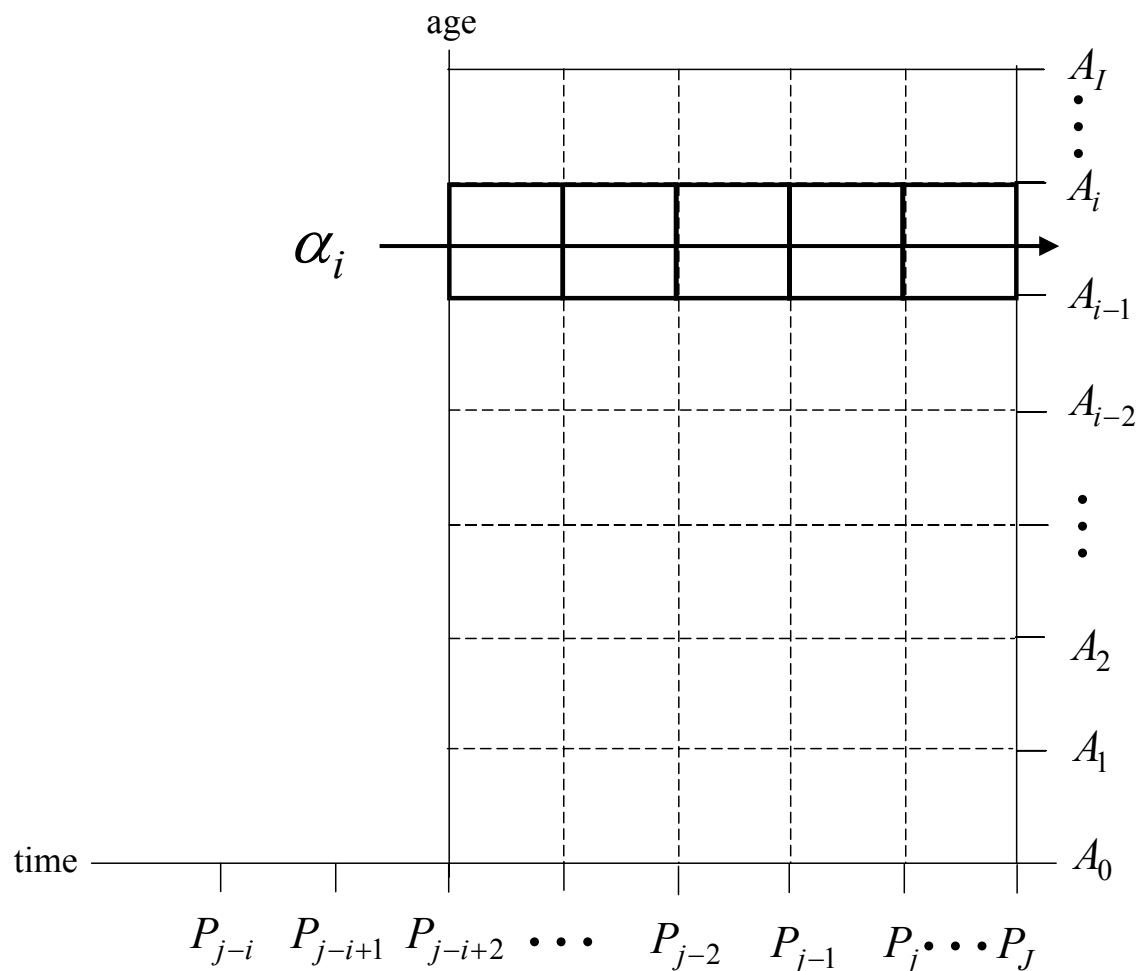
β_j : period effect

γ_{j-i} : cohort effect

Glenn (1977), Fienberg and Mason (1978), Osmond and Gardener (1982), Holford (1983), Kupper and Janis (1985), Clayton and Schiffllers (1987), Tango and Kurashina (1987)

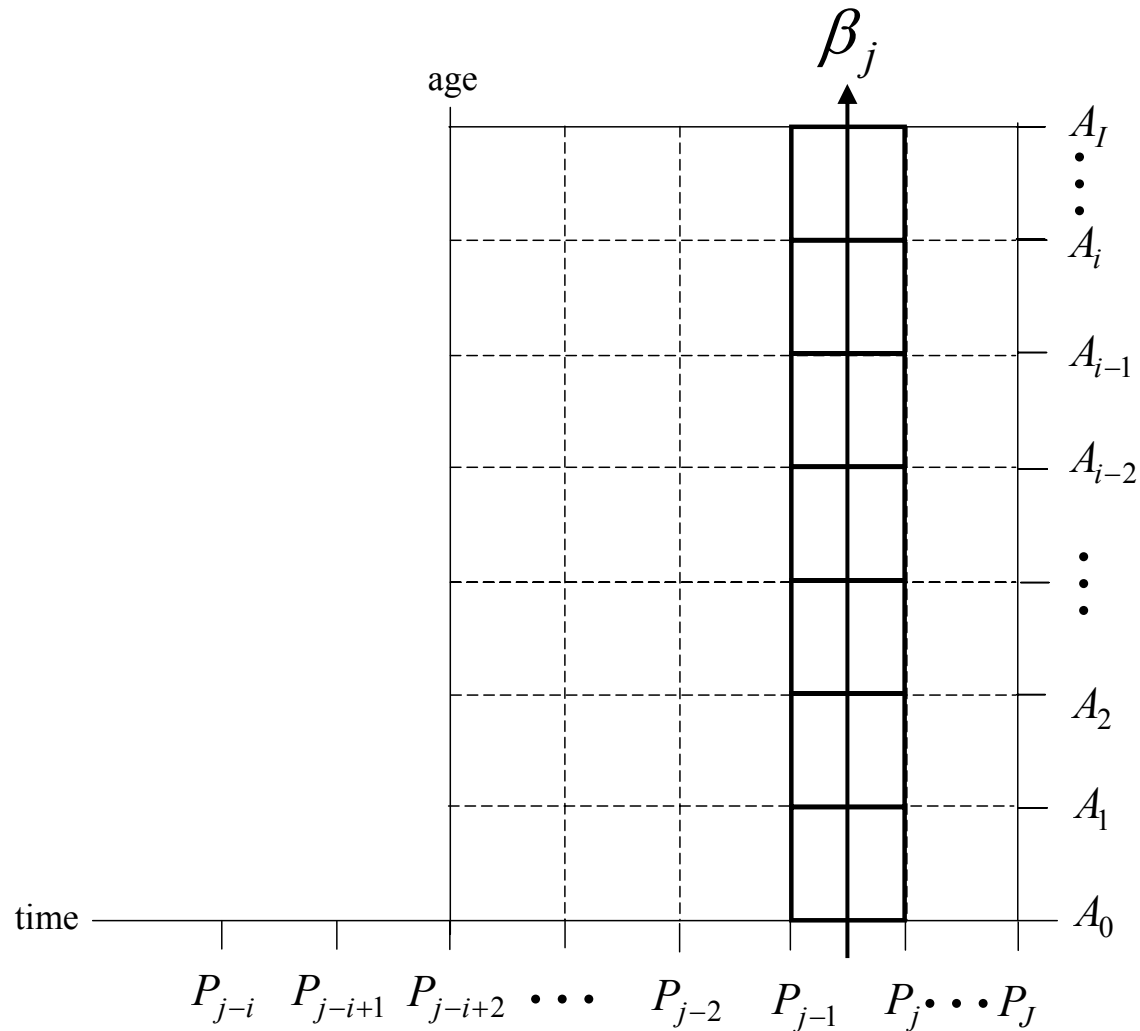
Age effect

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Period effect

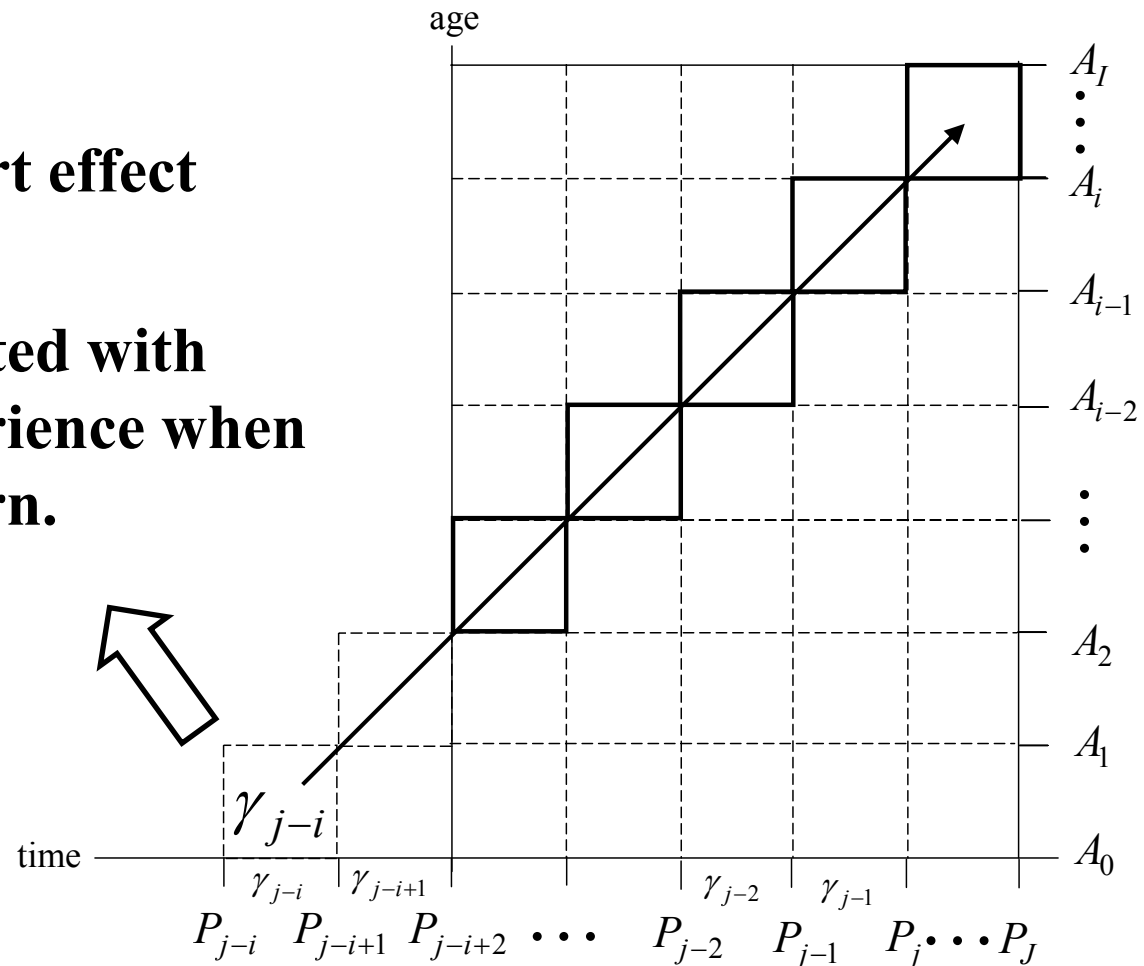
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Cohort effect

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Birth cohort effect
 ↑
**Effect associated with
 people's experience when
 they were born.**



Questions about birth cohort effect

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(i) Human experiences in their babyhood may be the most significant for determining their features.

Are experiences in other stages of life also significant?

(ii) It is supposed that the effect of experience in their babyhood appear immediately after they experience them.

Is such a effect equal to the period effect?

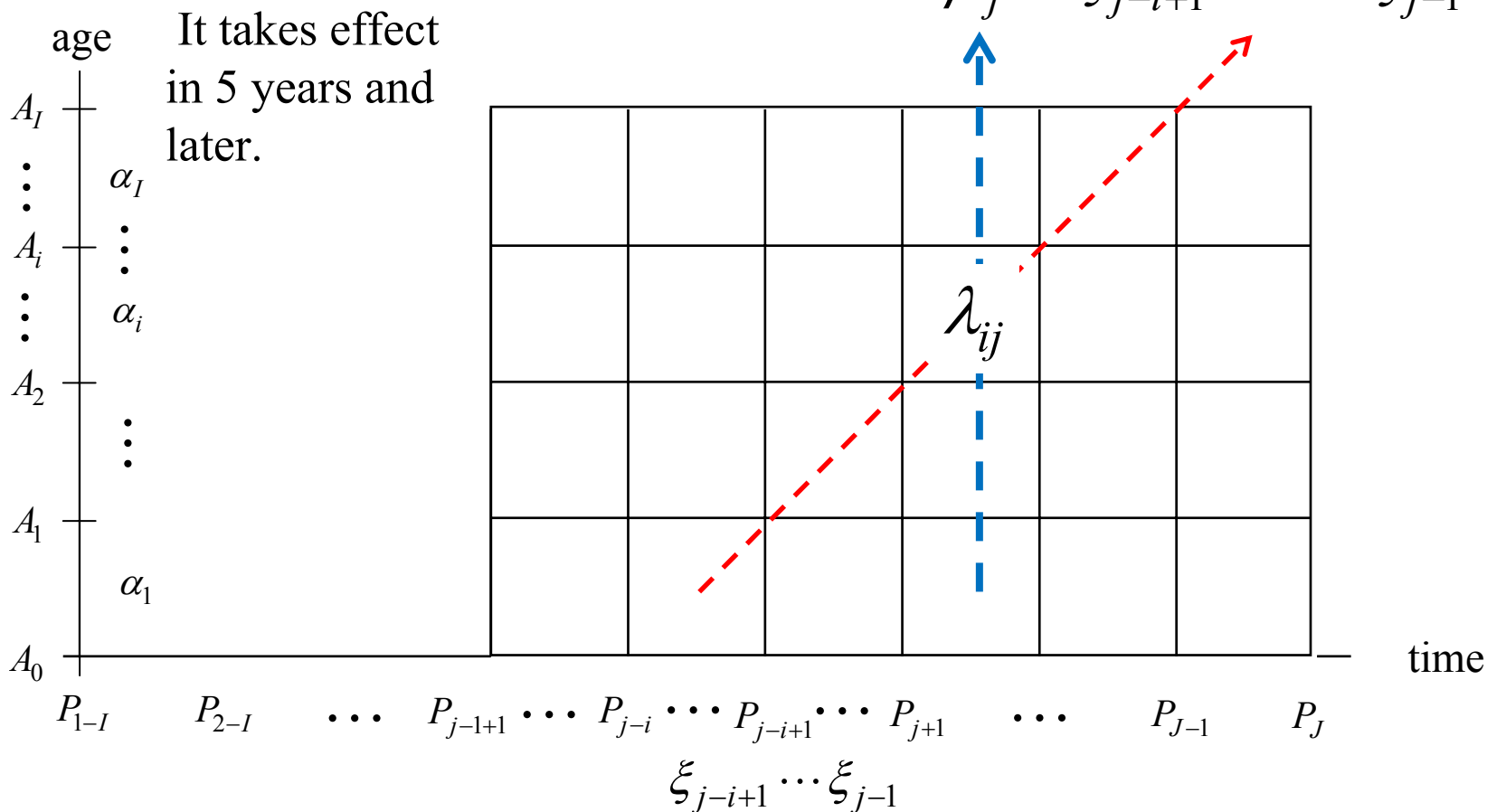


Are period and cohort effect overlapped
in their definitions?

Environment effect

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ξ_k : effect associated with exposure to the environment in the period $[P_{k-1}, P_k)$.



Age period environment model

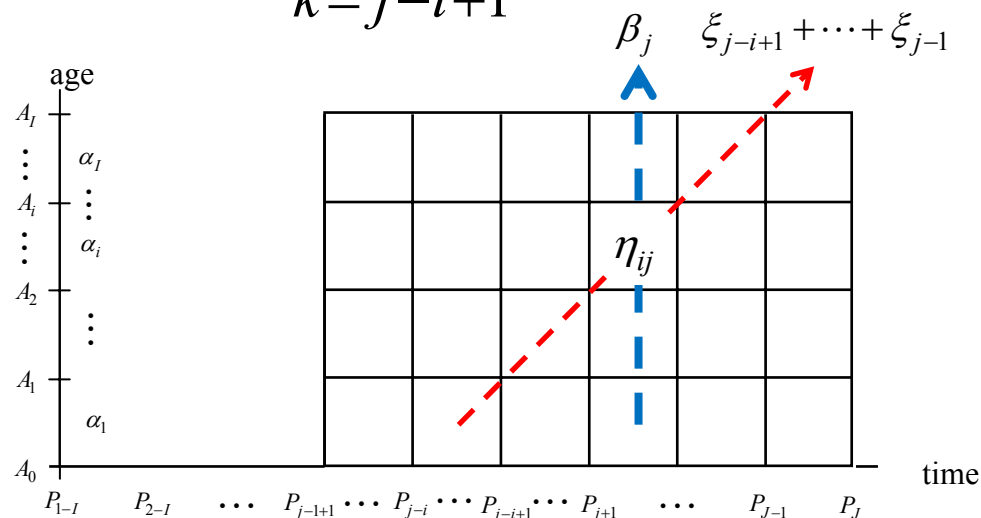
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$$\log \lambda_{ij} = \mu + \alpha_i + \beta_j + \sum_{k=j-i+1}^{j-1} \xi_k$$

α_i : age effect

β_j : period effect

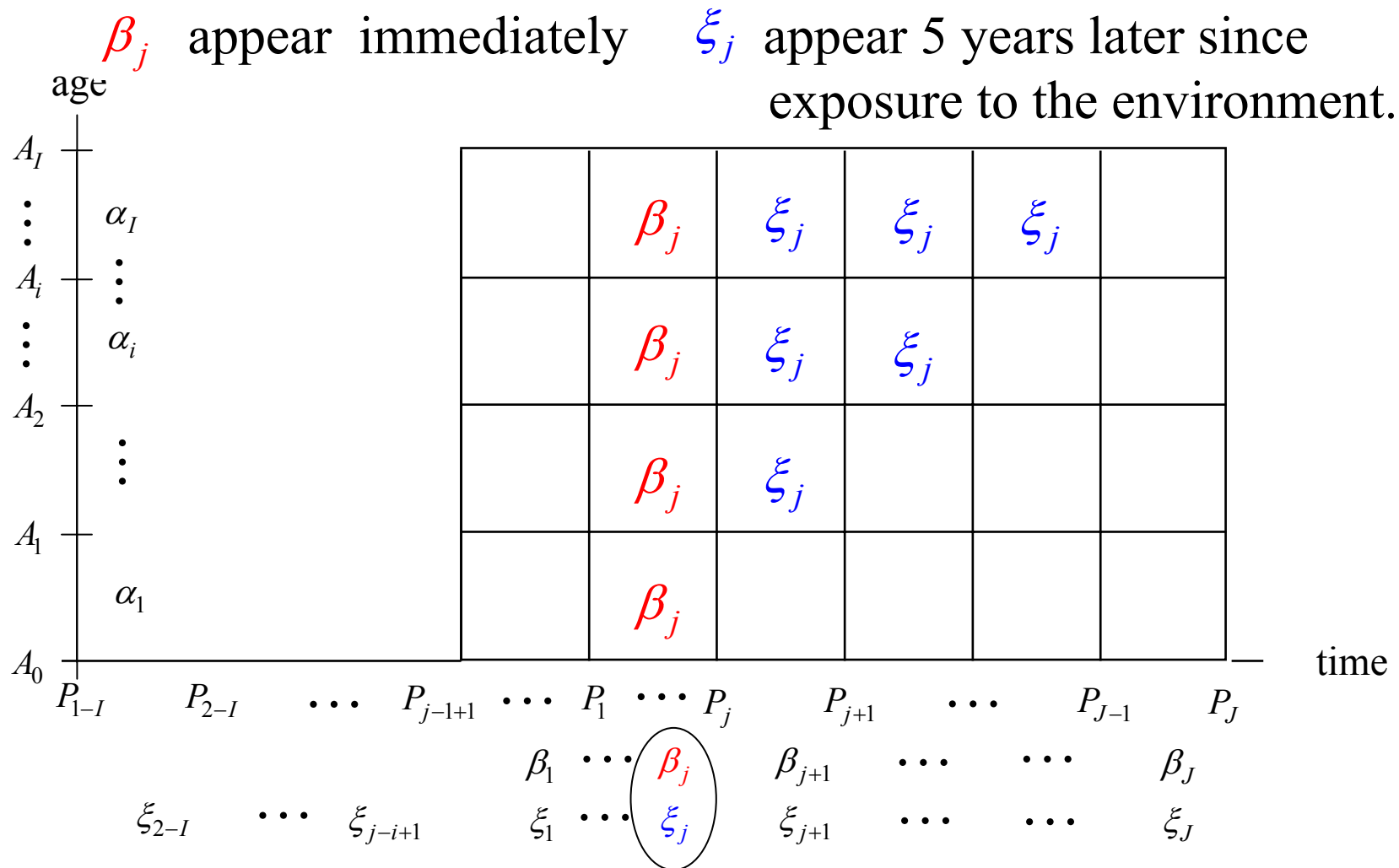
ξ_k : environment effect



Environment effects from people's birth to 5 years before present period are accumulated.

Period and environment effect

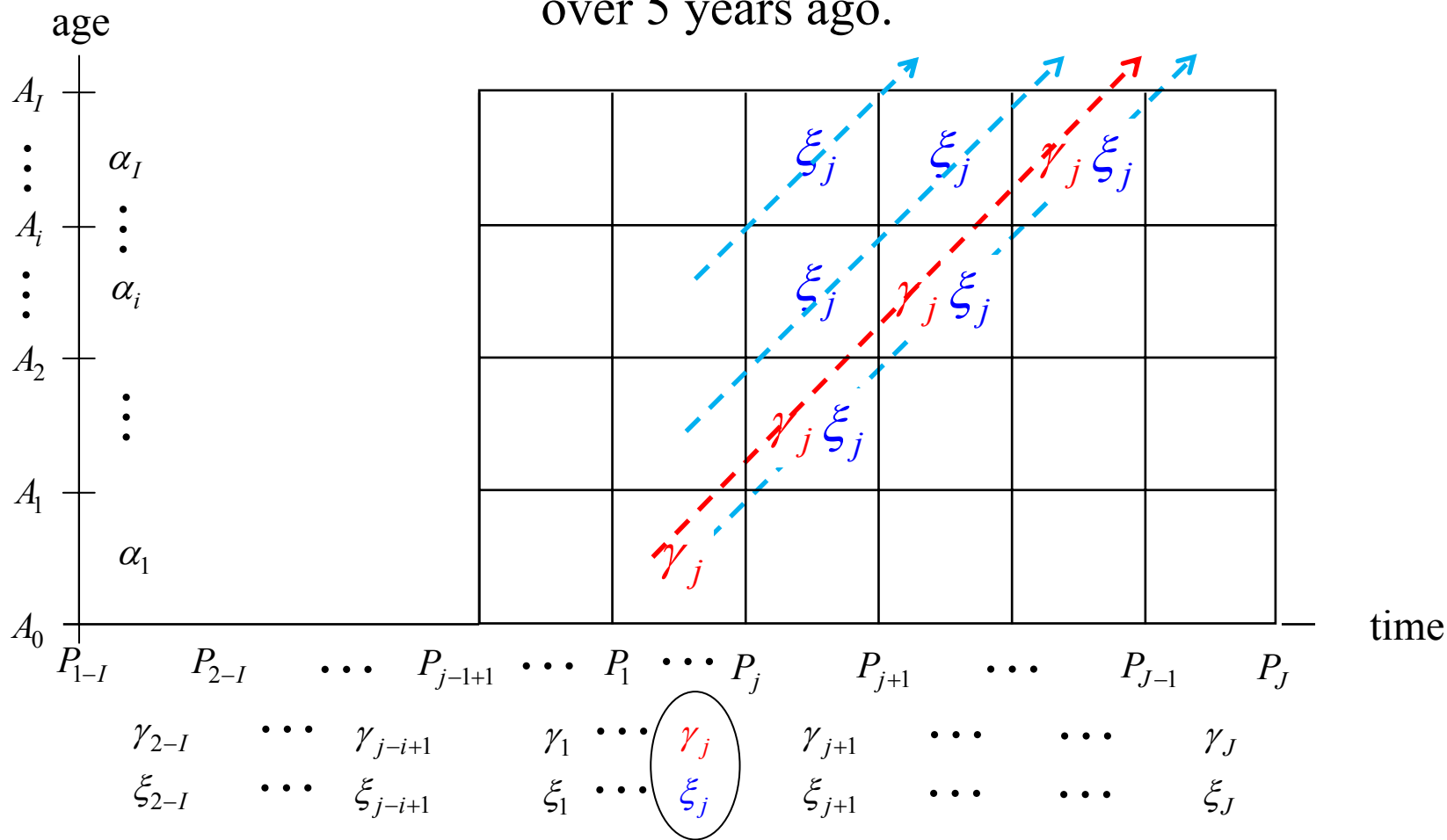
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Period and environment effect

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γ_j : affects one cohort ξ_j affects all cohort exposed to the environment over 5 years ago.

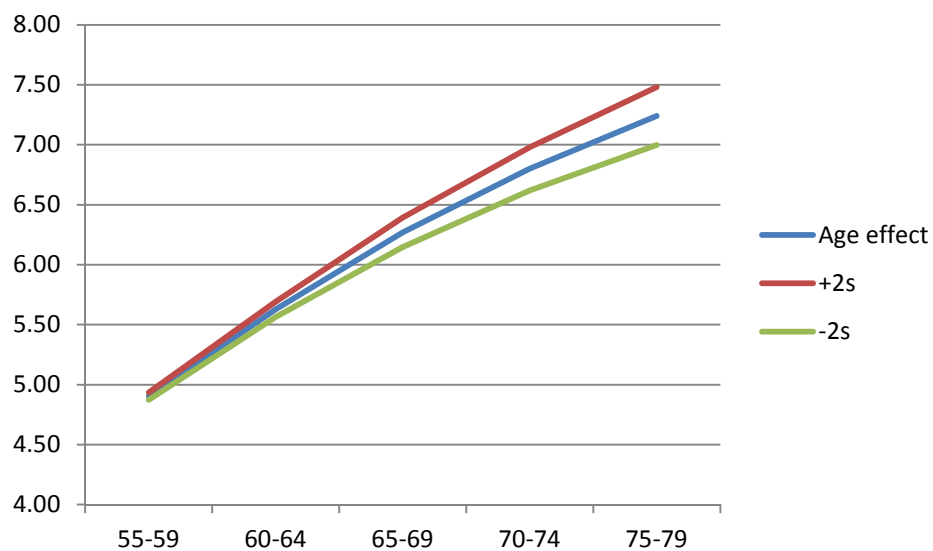


Results

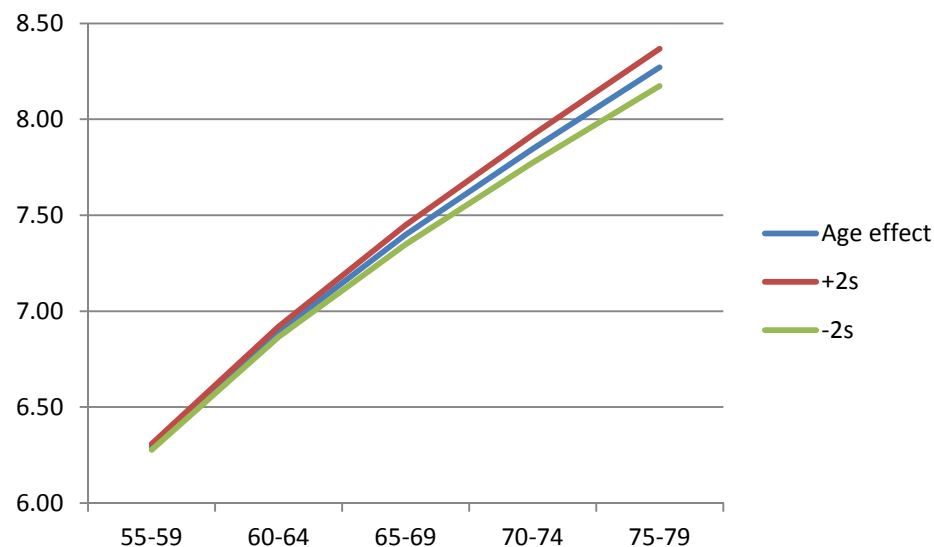
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Age effect on liver cancer

Female



Male

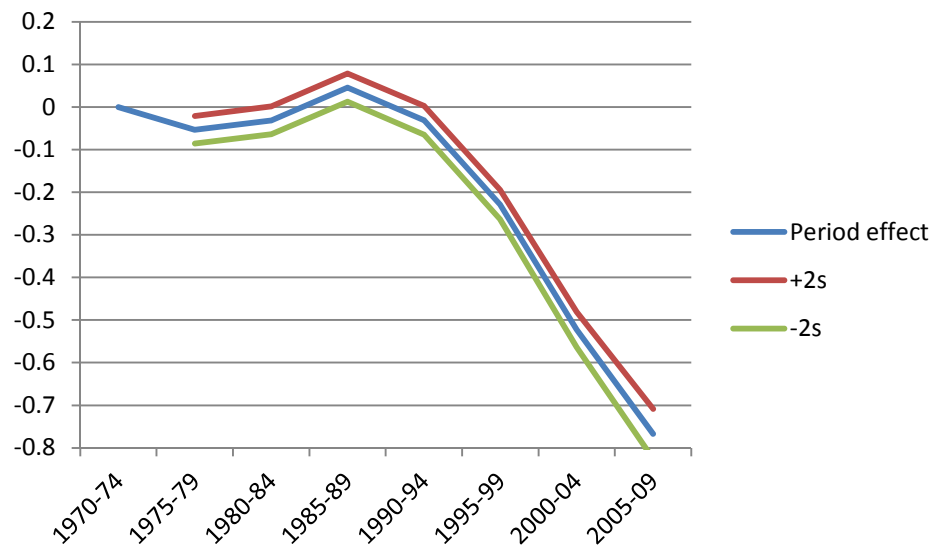


Results

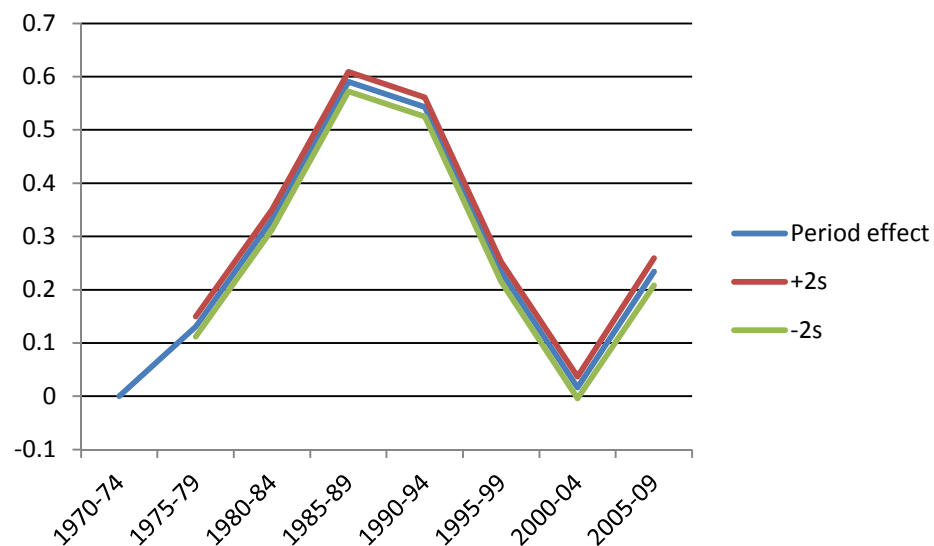
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Period effect on liver cancer

Female



Male

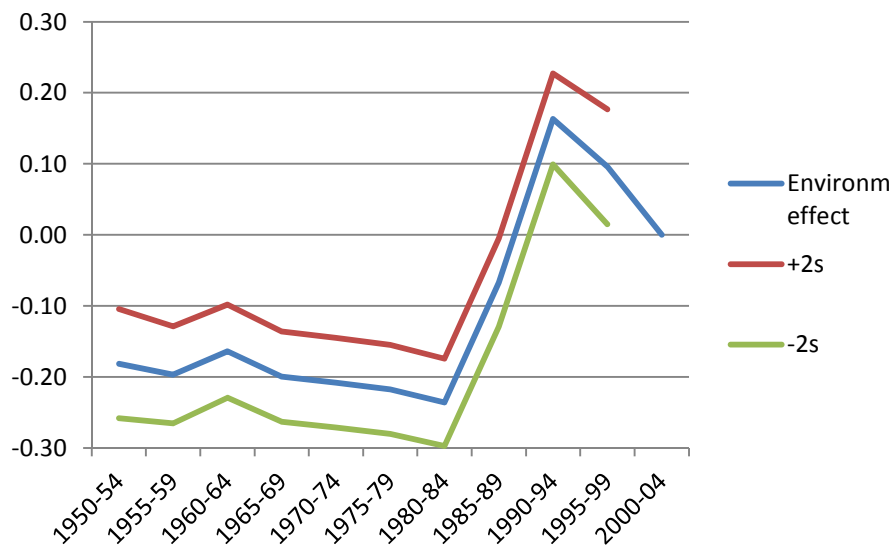


Results

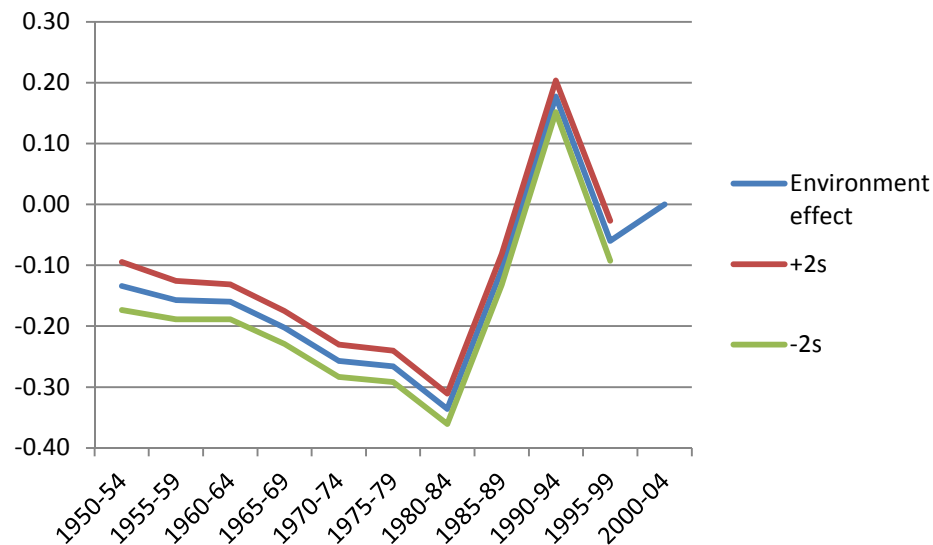
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Environment effect on liver cancer

Female



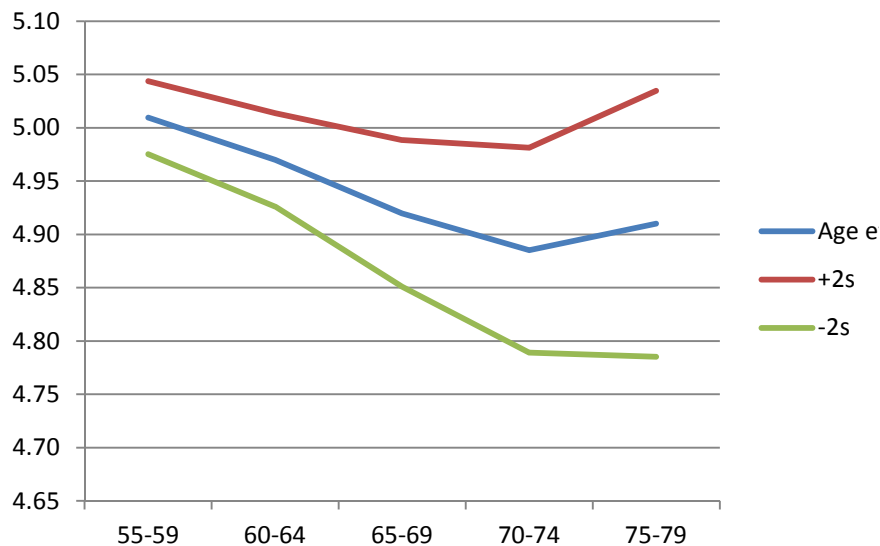
Male



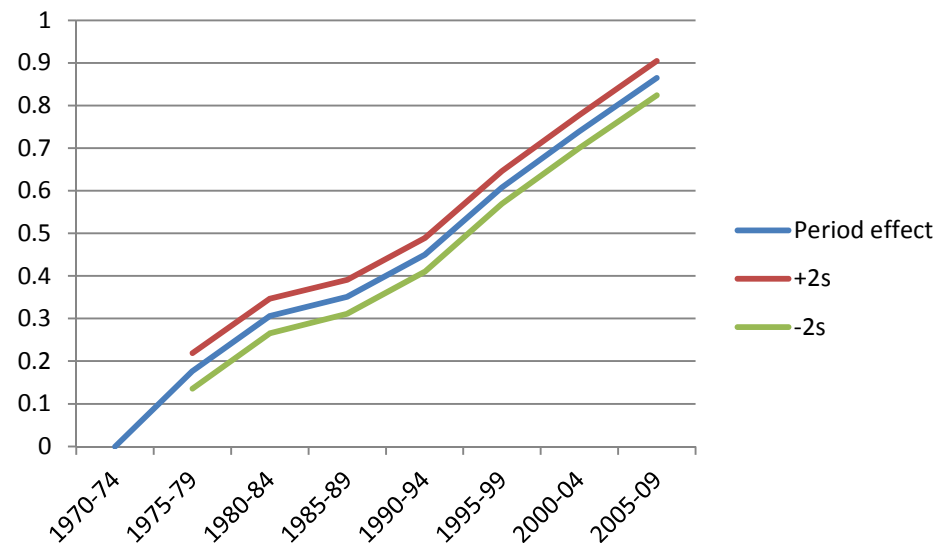
Results

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Age effect on breast cancer



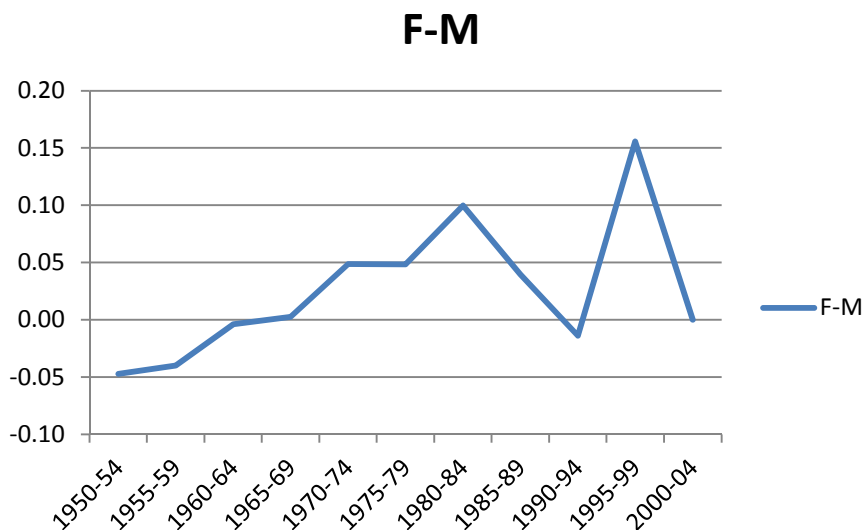
Period effect on breast cancer



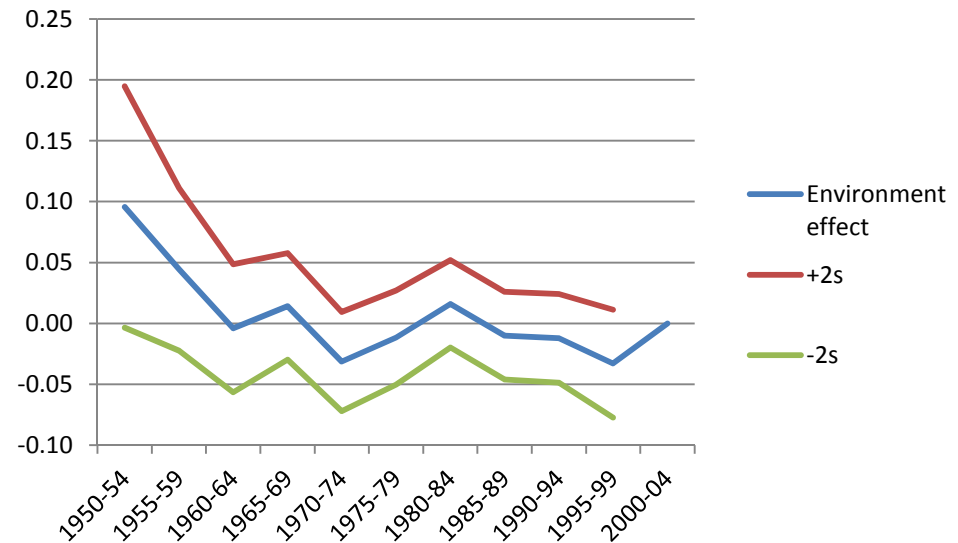
Results

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Difference of environment effect on liver cancer between sexes

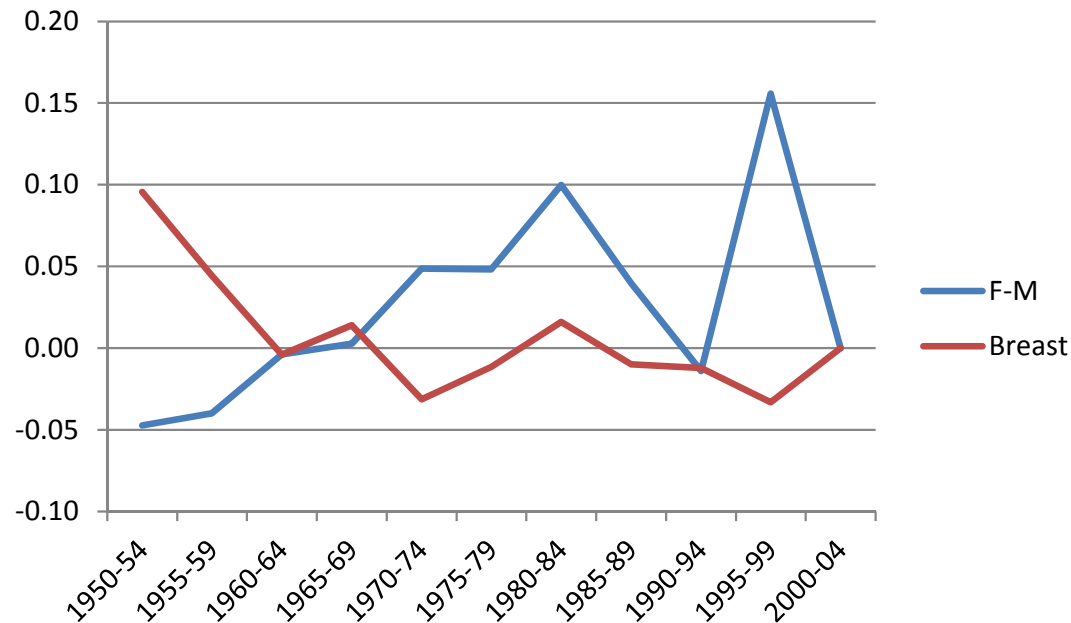


Environment effect on breast cancer



Results

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Those two factors have opposite trends.

Correlation Coefficient	t Stat	P-value
-0.62	-2.37	0.04

Conclusion

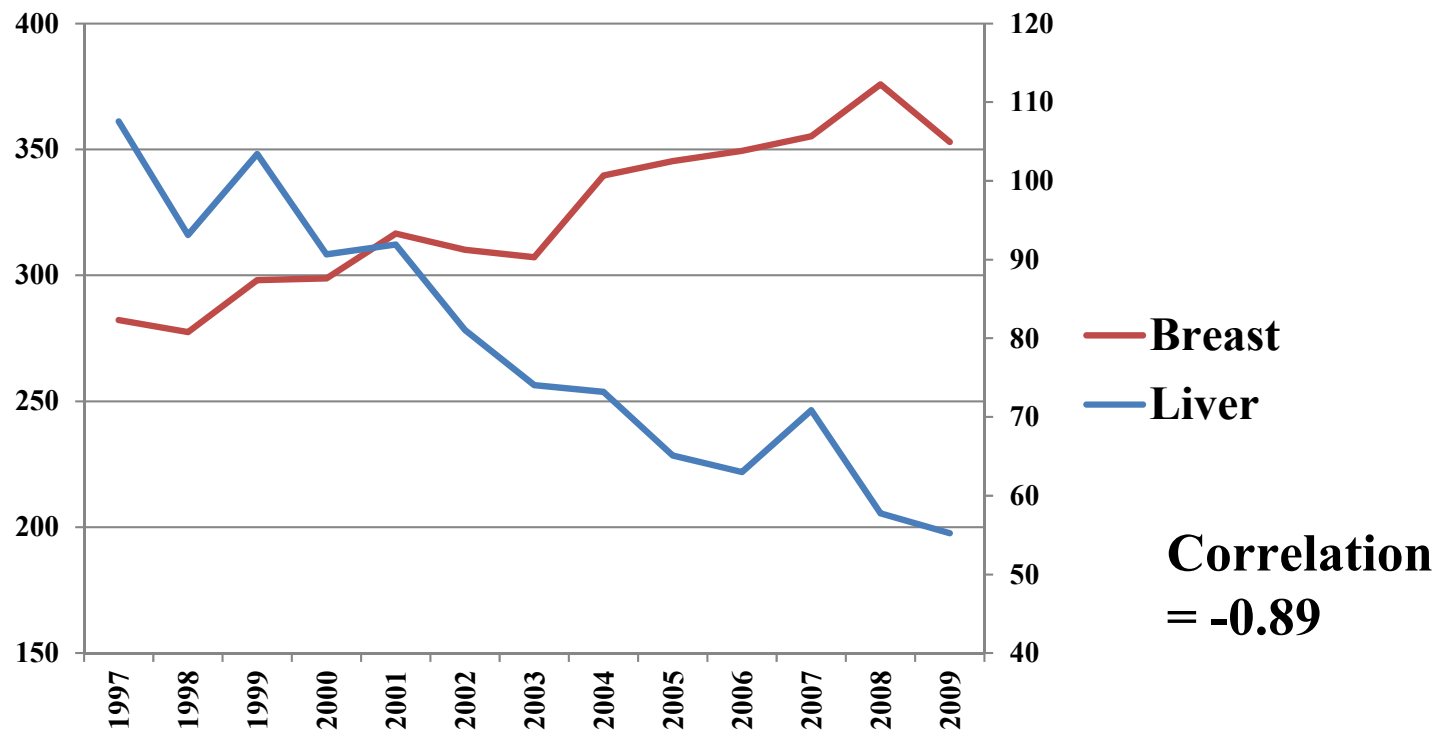
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It has been shown that there exist environmental factors which decrease the breast cancer risk and increase the liver cancer risk for postmenopausal females aged 55 and over.

Background

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Death rates of breast and liver cancer per 100 thousands Japanese women aged 55-59

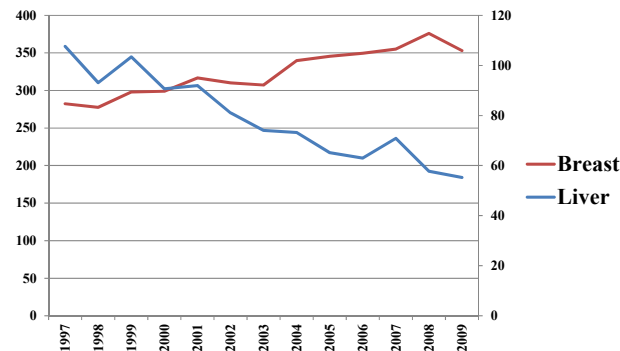


Background

The Joint Meeting of the 2011 Taipei International Statistical Symposium and IASC-CAR (Joint2011)

There exist environmental factors which decrease the breast cancer risk and increase the liver cancer risk for postmenopausal females.

Death rates of breast and liver cancer per 100 thousands Japanese women aged 55-59



There is some doubt that it is just a quasi-correlation?

Age period cohort model

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$$\log \lambda_{ij} = \mu + \alpha_i + \beta_j + \gamma_{j-1+1}$$

$$\begin{pmatrix} \eta_{11} \\ \eta_{12} \\ \eta_{13} \\ \eta_{21} \\ \eta_{22} \\ \eta_{23} \\ \eta_{31} \\ \eta_{32} \\ \eta_{33} \\ \eta_{41} \\ \eta_{42} \\ \eta_{43} \\ \eta_{51} \\ \eta_{52} \\ \eta_{53} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \\ \hline 1 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ \hline 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ \hline 1 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ \hline 1 & 0 & 0 & 0 & 0 & 1 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \\ \alpha_4 \\ \alpha_5 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \xi_{-3} \\ \xi_{-2} \\ \xi_{-1} \\ \xi_0 \\ \xi_1 \\ \xi_2 \end{pmatrix} .$$

Age period environment model

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$$\begin{pmatrix} \eta_{11} \\ \eta_{12} \\ \eta_{13} \\ \eta_{21} \\ \eta_{22} \\ \eta_{23} \\ \eta_{31} \\ \eta_{32} \\ \eta_{33} \\ \eta_{41} \\ \eta_{42} \\ \eta_{43} \\ \eta_{51} \\ \eta_{52} \\ \eta_{53} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 1 & 0 & 0 \\ 1 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 1 & 0 \\ 1 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 1 & 1 \\ 1 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 0 & 1 & 1 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 1 & 1 & 0 \\ 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 1 & 1 \\ 1 & 0 & 0 & 0 & 0 & 1 & 1 & 0 & 0 & 1 & 1 & 1 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 1 & 1 & 1 & 1 & 0 \\ 1 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 & 1 & 1 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \alpha_3 \\ \alpha_4 \\ \alpha_5 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \xi_{-3} \\ \xi_{-2} \\ \xi_{-1} \\ \xi_0 \\ \xi_1 \\ \xi_2 \end{pmatrix} .$$

The age-period-environment model is equivalent to APC model in terms of space spanned by column vectors in their design matrix.

Equivalency of the two models

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Correspondence between parameters

APC \leftrightarrow APE

$$\beta_j^{APC} = \beta_j^{APE} - \sum_{k=j}^{J-1} \xi_k \quad ; \quad j = 1, \dots, J-1$$

$$\beta_J^{APC} = \beta_J^{APE}$$

$$\gamma_k = \xi_k + \dots + \xi_{J-1} \quad ; \quad k = 2-I, \dots, J-1$$

APE \leftrightarrow APC

$$\beta_j^{APE} = \beta_j^{APC} + \gamma_j \quad ; \quad j = 1, \dots, J$$

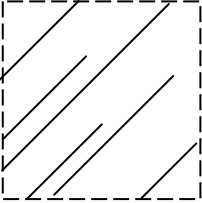
$$\xi_k = \gamma_k - \gamma_{k+1} \quad ; \quad k = 2-I, \dots, J-1$$

Lexis diagram

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The likelihood function for models assumed on (age period)-tabulated data is proportional to

$$L \equiv \prod_{i=1}^I \prod_{j=1}^J \lambda_{ij}^{y_{ij}} \exp(-\lambda_{ij} \cdot N_{ij})$$

X_{ij}  is approximated by person-years N_{ij} .
 $[A_{i-1}, A_i) \times [P_{j-1}, P_j)$

The likelihood estimation results in that for Poisson regression models.

Identifiability problem

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The likelihood function is a function only through λ_{ij} .

Therefore, it is needed to examine that for every set of

$$\{\lambda_{11}, \dots, \lambda_{IJ}\} \in \Lambda$$

the values of

$$\{\mu, \alpha_1, \dots, \alpha_I, \beta_1, \dots, \beta_J, \gamma_{1-I}, \dots, \gamma_{J-1}, \delta_2, \dots, \delta_I\} \in \Theta$$

are uniquely determined in the model.

If not, plausible **constraints** on the parameters
are needed.

Identifiability problem

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For any given μ , α_i , β_j and γ_{j-i} ($\alpha_1 = \beta_1 = \xi_{2-I} = 0$)

$$\mu + \alpha_i + \beta_j + \gamma_{j-i} = \mu + \alpha_i^*(d) + \beta_j^*(d) + \gamma_{j-i}^*(d)$$

holds for any d , where

$$\alpha_i^*(d) = \alpha_i + (i-1)d \quad \beta_j^*(d) = \beta_j - (j-1)d$$

$$\gamma_{j-i}^*(d) = \gamma_{j-i} - (j-i)d$$



**Linear trends in age, period and cohort effects
are not estimable.**

Identifiability problem

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In the proposed model under the constraint

$$\alpha_1 = \beta_1 = \xi_{2-I} = 0$$

the parameters are uniquely determined for every set of λ_{ij} .



The age, period and environment model are free from the non-identifiability problem from which APC model suffers.